

Risk-Based Dynamic Asset Allocation with Extreme Tails and Correlations

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Abstract

We propose a unique dynamic portfolio construction framework that improves portfolio performance by adjusting asset allocation in accordance with a forecast of market risk. We find that modifying asset allocation according to our market risk barometer offers investors the promising opportunity to meaningfully enhance portfolio performance across market environments.

Risk-Based Dynamic Asset Allocation with Extreme Tails and Correlations

Portfolio management is moving toward a more flexible approach capable of capturing dynamics in risk and return expectations across an array of asset classes [Li and Sullivan 2011]. The change is being driven, in part, by the observation that risk premiums vary as investors' cycle between risk aversion and risk adoration and that the decision to invest—whether to take risk and how much—is the most important investment decision [Xiong and Idzorek 2010]. Certainly, managers should take risks, but only if the returns appear to represent fair compensation. This all suggests that the traditional strategic approach of fixed-asset allocation is outmoded. The challenge of portfolio choice is much more than merely selecting for inclusion uncorrelated asset classes that constitute significant economic exposure and then specifying a fixed proportion of each.¹

Our effort facilitates this much needed dynamic flexibility to the asset allocation process. We propose a model of portfolio selection with heavy tails and dynamic return correlations. The powerful intuition behind our approach is that proper portfolio construction is an ongoing, dynamic process that integrates time-varying risks of the various asset classes within the investor's portfolio. We develop a dynamic asset allocation framework that determines an investor's optimal portfolio in accordance with changing global market environments and market conditions. Specifically, we consider how global return, variance, and covariance characteristics vary across time and states of global markets for a diversified portfolio of asset classes. We then use this dynamic information to consider the asset allocation implications in a practical setting. Our novel approach builds on the regime switching framework of Ang and Bekaert (2002, 2004), Kritzman and Page (2011), among others, and provides a framework that illuminates the changing nature of global market risks and directs accordingly asset allocation and risk decisions.

We argue that it is imperative for managers to monitor and react to changes in the macro-environment on an ongoing basis. Our effort provides one such useful framework— a genuine barometer for monitoring risk dynamics across our global financial system and reacting to those market conditions across time.

¹ For further discussion on this topic, see Sullivan (2008)

The Framework for Dynamic Asset Allocation

The framework we offer has important implications for portfolio risk management and asset allocation decisions. It takes into account skewness and kurtosis, moments of the return distribution beyond mean and variance, as well as persistence in volatility, or volatility clustering, and correlations of risky asset returns which tend to increase during times of market turbulence, or return dependence. Our non-linear model framework is more dynamic and less restrictive than traditional, static methods that depend on returns following a Gaussian process. One practical application of our approach is that it provides a monitoring device regarding market instability and portfolio vulnerability. Furthermore, we demonstrate that investors can act before the iceberg is under the ship's keel. The result is a high frequency, dynamic technique that allows investors to proactively monitor and manage portfolio risk via real-time asset allocation decisions.

We dynamically and proactively determine asset weightings as conditioned on changing market volatility and covariances. Asset allocation is further accomplished in accordance with one of two possible states of the world: normal risk (normal uncertainty: normal return volatility and correlations), and high risk (high uncertainty: high volatility and correlations). Behind the two states lies a mechanism driven by factors determined to possess predictive power of the degree of economic and market uncertainty governed by *forward* transition probabilities where the regime variables are used to fit a Markov Regime Switching process [see Ang Bekaert 2002, 2004]. Our regimes correspond to market dynamics and the non-normal return distributions characterizing markets [e.g., Xiong and Idzorek 2011, and Sullivan, Peterson, Waltenbaugh 2010]. We do not model changes in expected returns, which are known to be particularly difficult and often leads to models biased by hindsight and model over fitting.

At a high level, the strategy we propose consists of three main, overarching parts. In the first part, we estimate the conditional value at risk (CVaR) for a market representative portfolio [Kaya, Lee, Pornrojngkool 2009]. The estimated CVaR then serves as critical input into our second part, a forecast of market risk—modeling the probability that markets are in, or about to enter a turbulent financial period. This

information then enables the third part— proactively adjusting the portfolio asset allocation in accordance with the market risk regime forecast obtained part 2.

We begin by applying extreme value theory (EVT)² which allows us to model fat-tailed return distributions for a host of asset classes with particular attention to volatility clustering and extreme co-movements across various markets [e.g., Sullivan, Peterson, Waltenbaugh, 2010]. The asset classes included in our framework are: global equity, U.S. investment grade bonds, U.S. high yield bonds, commodities, and U.S. real estate investment trusts. Our base case portfolio asset allocation, described in Exhibit 5, is constructed based on weights typically found in institutional portfolios, and close to the capital market weights.

We employ conditional value at risk (CVaR)³ to facilitate forward looking scenario-based outcomes outside the range of historical observations. A two-state Markov-Switching model is applied to identify regimes in the forward-looking market downside risk measure, CVaR. The CVaR then forms the basis for our dynamic risk and asset allocation framework by providing an indicator of downside risk across markets and for optimization in portfolio construction.

Altogether, we build an effective regime-dependent investment strategy based on market downside risk and asset class co-movements across time. To accomplish this task, we follow a dynamic asset allocation framework under a Mean-CVaR optimization approach with varying target CVaR according to market regimes. The end result is an implementable tail risk management process in accordance with the increasingly interconnected and dynamic risks observed in markets.

Data and Model Setup

² Readers are referred to Embrechts, Klüppelberg, and Mikosch [1997] for a comprehensive treatment of extreme value theory.

³ CVaR measures the expected loss during a given period at a certain confidence level. As a better alternative to VaR, it incorporates both the possibility and expected magnitude of loss. Moreover, it is coherent and convex and can readily be incorporated into discrete optimization process in risk management [Uryasev 2000 and Rockafellar 2002]. For example, a 95% 21-day CVaR of 20% means the investor expects to lose 20% within the 5% worst-case scenarios in a month. CVaR is known as mean excess loss for continuous distributions, and defined as the weighted average of VaR and losses strictly exceeding VaR for discrete distributions.

Exhibit 1 provides an overview of the five asset classes included in our analysis along with summary statistics. All asset classes are represented by indexes in the following way: global equities by the Morgan Stanley Capital International ACWI Index (MSCI ACWI), commodities by the Goldman Sachs Commodity Index (SPGSCI) total return index, U.S. real estate by the Dow-Jones Wilshire REIT (DW REIT) total return index, U.S. high yields bonds by Merrill Lynch High Yield Master II (MLHY II) total return index, and U.S. investment grade bonds by the Barclays Capital Aggregate Bond Index (Barclays Agg.) gross return index. All summary statistics are based on daily data (not annualized) from February 1, 1996 to October 10, 2011. In reviewing Exhibit 1, we draw the reader's attention to the negative skewness observed for almost every asset class (except REITS), and the excess kurtosis across all asset classes, especially for REITs and high yield bonds.

Exhibit 1

Consistent with prior research, further examination of the data reveal that autocorrelation is present in the return series, especially for day $t+1$. This can be seen visually for MSCI ACWI by the autocorrelation functions for the log of daily returns and the square of log returns, or variance, shown in Exhibit 2 Panel A. We return to address these issues which motivate our analysis, later.

Exhibit 2

Forecasting Market Risk

For the first phase in our three-part framework—a daily forecast of the risk of the overall portfolio—the model we employ for the joint fat-tailed distribution of returns and the subsequent calculation of CVaR involves the 5 main steps outlined below.

- 1) *Return Filtering.* We filter each daily return series using AR(1)/GJR-GARCH(1,1) process to remove serial correlation and standardize the residuals;
- 2) *Marginal Distribution Modeling.* We employ a peaks-over-threshold method to estimate the marginal semi-parametric empirical CDF of the filtered standardized residuals from step 1 [e.g., Focardi, Fabozzi 2004, Tsay 2005]. We use a non-parametric Gaussian kernel to derive the interior portion of the distribution and a parametric GPD to estimate the left hand and right hand tails;
- 3) *Extremal Dependence Modeling.* We transform the standardized residuals from step 1 into uniform variates using the semi-parametric empirical CDF derived in step 2. We then fit a t-copula to the transformed data to allow for joint “fat” tails.
- 4) *Return Simulation.* Given the parameters of the t-copula, we simulate 21 dependent uniform variates for all indices 10,000 times. Then via the inversion of the semi-parametric marginal CDF for each index, we transform the uniform variates to standardized residuals that are independent in time but dependent at any point in time. Last, we reintroduce the autocorrelation and volatility clustering observed in the original index using parameters obtained from step 1 to arrive at the simulated 21-day daily returns for all five asset classes.
- 5) *Risk Forecasting.* We forecast 21-day market representative portfolio risk with the the policy allocation as shown in Exhibit 5 serving as the baseline. The average 21-day portfolio loss in the worst 5% scenarios based on the 10,000 simulations becomes the portfolio 95% CVaR. This CVaR is then used as the across market tail risk indicator in the second part of our three-part framework—regime dependent dynamic asset allocation. Expected returns are also shown in Exhibit 5, and do not change for any regime environment.

We now discuss in more detail the five steps outlined above used to arrive at our dynamic, high-frequency estimate of portfolio risk using CVaR and extreme value theory (EVT). Modeling the tails of a distribution using EVT requires the observations to be approximately independent and identically distributed (i.i.d). As a consequence, we first filter our return series with the aim of the filtering process to produce approximately i.i.d observations. To accomplish this objective, for each return series we fit a first order autoregressive model AR(1) to the conditional mean of the daily log returns using equation (1) and an asymmetric GJR-GARCH(1,1) [Glosten, et al., 1993] to the conditional variance using equation (2), below.

$$r(t) = c + \theta r(t - 1) + \epsilon(t) \quad (1)$$

$$\sigma^2(t) = \kappa + \alpha \sigma^2(t - 1) + \phi \epsilon^2(t - 1) + \psi [\epsilon(t - 1) < 0] \epsilon^2(t - 1) \quad (2)$$

$$z(t) = \epsilon(t)/\sigma(t) \text{ i.i.d distributed } t(v)$$

With this model, we address the so-called leverage effect whereby a negative association has been observed to exist between shocks to asset returns and future volatility [Black 1972]. Specifically, the last term of equation (2) incorporates asymmetry into the variance through the use of a binary indicator that takes the value of 1 which predicts a higher volatility for the subsequent day if the prior residual return is negative, and a takes on a value of 0 otherwise. We then standardize the residuals by the corresponding conditional standard deviation as commonly done for such exercises. Finally, the standardized residuals are modeled using the standardized Student's t-distribution in order to capture the well-known fat tails in the distribution of returns.

The result of this process is shown in Exhibit 2B which plots the autocorrelations of the standardized residuals for the MSCI ACWI return series. As seen from Exhibit 2B, the filtering process we employ results in approximately i.i.d. observations and thus volatility clustering has been eliminated by the filtering process. The resulting standardized residual returns approximate a zero-mean, unit-variance, i.i.d series. This allows us to employ EVT estimation of the tails from our sample cumulative distribution function (CDF).

As EVT allows only for estimation of the tails of the distribution, we combine these tail distributions with a model for the remaining internal part of the distribution. To accomplish this task, we move to step 2 and follow the peaks-over- thresholds approach [McNeil 1997] and define upper and lower thresholds as that set of minimum residual returns (we use the 90th percentile) found each of the left hand and right hand tails. The result is a partition of the standardized residuals into three regions; the lower tail, the interior, and the upper tail. A non-parametric Gaussian kernel CDF is used to estimate the interior of the distribution. We then fit those extreme residuals in each tail beyond the thresholds using EVT. In particular, we use a parametric Generalized Pareto Distribution (GPD) estimated by maximum likelihood. The CDF of the GPD is parameterized using equation 3, with exceedances (y), tail index parameter (ζ) and scale parameter (β).

$$F(y) = 1 - \left(1 + \frac{\zeta y}{\beta}\right)^{-1/\zeta}, y \geq 0, \beta > 0, \zeta > -0.5 \quad (3)$$

Exhibit 3 shows a visual representation of the upper and lower tails of the return distribution for ACWI. It shows that our GPD approach far better accommodates the fat tails observed historically in the return distribution. As can be seen from Exhibit 3, the GPD curve much more closely approximates the historical, or empirical, return distribution, and as such, allows for a more accurate representation of the reality of fat-tails.

Exhibit 3

With our fat-tailed conditional distribution of returns in place, we can now turn attention to the next important element in risk modeling, step 3— how asset class returns move together in the extremes. For our extremal dependence model, we consider asset return covariances via the joint distribution of returns using copula theory (Focardi, Fabozzi 2004). With copulas, we are able to model the observed increased co-dependence of asset class returns during periods of high market volatility and stress. Empirically, not only do individual asset classes have “fatter” tails than that allowed in a normal, Gaussian distribution, combinations of asset classes also

exhibit a higher incidence of joint negative returns in times of market stress. That is, risky asset returns across asset classes abruptly decline in unison. By way of example, as shown in Exhibit 4, both MSCI ACWI and GSCI have occasionally realized simultaneous loss events of four standard deviations or more. A bivariate normal distribution would therefore provide a poor representation of the dynamics of these joint jumps observed in asset class returns observed in recent decades. A more realistic approach is needed.

EXHIBIT 4

To account for the incidence of returns abruptly moving in unison, we employ copula theory which accommodates interrelated and extreme dependencies of returns. More specifically, copulas allow for the modeling of fat tails even when asset class returns present a high degree of co-movement as seen historically. We chose to employ the t-copula because this particular copula enables us to better capture the effects of fat tails and allocate non-zero probabilities to observations which may occur outside of the range of historical returns. By adjusting the copula's degree-of-freedom parameter, we can extrapolate our multivariate fat-tailed distributions so that it is consistent with the observed empirical data. Having estimated the three regions of each marginal semi-parametric empirical CDF, we transform them to uniform variates, and then fit the t- copula to the transformed data.

We can now move to step 4 and generate our scenario-based forward looking projections of downside risk across markets using Monte Carlo simulations. Given the parameters of the t-copula from step 3, we simulate 21 dependent uniform variates of all five indices 10,000 times. Then via the inversion of the semi-parametric marginal CDF of each index, we transform the uniform variates to standardized residuals to be consistent with those obtained from the AR(1)/GJR-GARCH(1,1) filtering process in step 1. These residuals are independent in time but dependent at any given point in time. Here, we reintroduce the autocorrelation and volatility clustering observed in

the historical returns for each index. This allows us to move to step 5 whereby we aggregate the portfolio and project a 21 forward day downside risk for the aggregate portfolio. This downside risk is measured as the 95% CVaR, and is the average portfolio loss in the worst 5% scenarios, based on 10,000 Monte Carlo simulations.

To generate the time series of our 21 day look-ahead portfolio risk forecast, we repeat the steps above and forecast the portfolio 95% CVaR under an expanding window approach. To avoid look-ahead bias, we incorporate only that market information available at the time the model forecast is generated. The result of our risk forecast effort is shown in Exhibit 6 as represented by our 21-day forward combined portfolio CVaR for the base portfolio. As can be seen from Exhibit 6, our portfolio risk estimate is highly responsive to actual market dynamics.

EXHIBIT 5

EXHIBIT 6

Forecasting Market Risk Environments

In the next part of our framework, we estimate the probability that the market environment is already in or about to enter a turbulent state and use this information to inform our asset allocation decision. Here, our asset allocation is determined in accordance with one of two possible states of the world; normal risk (normal uncertainty: normal return volatility and correlations), and high risk (high uncertainty: high volatility and correlations, low returns). The two market states are governed by a *forward* transition probability forecast of CVaR derived earlier. Specifically, our CVaR forecast is used as the regime variable to fit a two-state Markov Regime Switching

process [see Ang Bekaert 2002, 2004]]. In this way, our regimes correspond to market dynamics and the non-normal return distributions characterizing markets

Exhibit 7 reveals the meaningful presence of a normal regime and an event regime in our time-series forecast of market downside risk. This is evidenced by the substantial change in both the mean and the standard deviation of our CVaR regime variable. Over the estimation period, the high-risk, event regime shows an average 21-day CVaR (95%) of -14.22% with a standard deviation of 5.69%, as compared to a higher average CVaR of -6.12% with a lower standard deviation of 1.63% for the normal regime.

EXHIBIT 7

In general, the Markov-Switching model we use seeks to more effectively capture the dynamic volatility of the regime variable as compared to simple data partitions based on arbitrary thresholds. To understand why this is so, consider that if the prior CVaR estimate suggests a high volatile (normal) state, the model would more likely predict that the current market environment is also a high volatile (normal) state. A naïve, fixed threshold may not make the same association and may thus classify the current state as part of the normal regime, if the current CVaR value is below the arbitrarily chosen threshold. In short, the regime model we employ is better equipped to adapt over time to changing market conditions in real time.

Exhibit 8 shows the time-series results of the resulting forecast of the probability that the markets are in, or about to be in, a high risk state (regime probability bigger than 50%) over time. To estimate our model, we use an expanding window approach with our first estimate in January 3, 2000 using data from February 1, 1996 to January 2, 2000. We generate each new forecast daily by simply adding new observations and re-estimating the model with the new observations as the data become available. The results, shown in Exhibit 8, Panel A, highlight that our Markov-Switching model succeeded in meaningfully partitioning the market into two

regimes. Exhibit 8, Panel B, shows the specific dates identified as the market being in a high-risk “event” regime defined as an event probability of at least 50%.

EXHIBIT 8

A further understanding of the impact of our regime risk model on asset class performance can be inferred from the data presented in Exhibit 9. Here, we summarize the risk and return statistics for each of our five asset classes during the study period, January 3, 2000 to October 10, 2011. A comparison of Exhibit 9A (event days) and Exhibit 9B (full period), shows that during the event periods the median returns for all risky assets are lower and standard deviation of returns are all higher, versus the full period. These results suggest that the model assisted in anticipating turbulent periods.⁴ Furthermore, extreme returns are shown to be a dominant presence during forecasted event regimes. This can be seen from the percentiles, e.g. the 5th percentile and 95th percentile are much further apart for the event regime daily return distributions versus the full period.

EXHIBIT 9

Dynamic Asset Allocation

We now discuss the third, and final, part of our modeling; incorporating our forecast of market turbulence into an effective dynamic asset allocation framework. Our portfolio construction process responds to market dynamics by adjusting the overall portfolio asset allocation in accordance with our regime-based risk forecast and

⁴ Furthermore, the maximum and minimum daily returns always occurred in the high volatile event regime suggesting that investors might benefit from a regime model that can correctly distinguish a third regime for high return periods.

Mean-CVaR optimization.⁵ As mentioned, we employ a risk-on (risk-off) approach as driven by our model prediction for either a normal or high risk state, respectively. Importantly, the dynamic portfolios we construct here facilitate a direct evaluation of the risk present in markets with an eye towards mitigating the impact of abrupt downside events frequenting markets via dynamic asset allocation.

Specifically, we solve equation 3 to obtain the weights of a portfolio that maximizes expected return while targeting the CVaR to a desired level (see Rockafellar and Uryasev [2002]). Expected returns and benchmark portfolio weights are shown in Exhibit 5. This approach allows us to incorporate our copula-driven fat-tailed simulation scenarios into a portfolio allocation optimization problem. Furthermore, as we will see, this allows the optimal portfolio allocation to be determined in accordance with our market regime prediction. Specifically, the fixed expected return vector is represented by μ , and w is the set of weights that belongs to the space X . We examine both an unconstrained portfolio with no-shorting and no-leverage portfolio (weights must be between 0% and 100%) and a constrained portfolio with bounds as shown in Exhibit 5. The CVaR target constraint is represented by $\tilde{\theta}_{95\%}(w)$ and is the resulting forward-looking CVaR at a 95% confidence level as estimated given the set of weights, w , with a target CVaR level of γ .

$$\begin{aligned} & \text{Maximize } \mu w \\ & \text{Subject to } \tilde{\theta}_{95\%}(w) \leq \gamma, w \in X \end{aligned} \quad (3)$$

We next demonstrate the approach by back-testing the model outcomes combining all three parts of our process over time. We explore both unconstrained and constrained portfolio weighting schemes as shown in Exhibit 5. As mentioned, we reduce the effect of any hindsight bias on our results by using static, unadjusted expected returns. The main focus of this paper is to show the meaningful impact that can be had on portfolio performance by adjusting “only” the portfolio asset allocation in accordance with dynamic forecasts of market risk as captured by changing variance and covariances across asset classes, over time. To this end, we forecast risk and

⁵ There is an extensive literature on advanced portfolio optimization techniques. See, for instance, Fabozzi, et al., 2007 and Rachev, Stoyanov and Fabozzi 2008.

rebalance the portfolio according to pre-specified rules discussed below. For all results, we use static expected returns and the policy portfolio as the benchmark portfolio, as shown in Exhibit 5.

Exhibit 10 shows the results from our portfolio construction process. In Panel A, we employ a set monthly rebalancing rule whereby we rebalance the portfolio every 21-days. We compare the performance of the benchmark portfolio to an unbounded (weights must be between 0% and 100%, e.g., no shorting and no leverage) portfolio construction processes, all rebalanced each 21 days. The two unbounded portfolios are optimized portfolios based on CVaR, as discussed above. We show results for several static target levels of CVaR, and we then allow the target level of CVaR to switch over time between a high- and low-risk level in accordance with our dynamic regime forecast.

In row 1 of Exhibit 10, Panel A, we show the performance of the benchmark portfolio. We compare outcomes for our unbounded portfolios which allow for the weights for each of our five asset classes to vary between 0% and 100% over the study period. First, we show the performance of overall portfolios created by imposing a series of constant, maximum allowable level of mean-CVaRs. Here, we report the results for five constant target CVaR levels ranging from lowest risk (3 percent CVaR) to highest risk (7 percent CVaR). This allows a comparison of how various CVaR limits reflect changing risk conditions as estimated solely by our CVaR model, while temporarily ignoring our market risk regime forecast in constructing portfolios. As expected, portfolio draw downs and volatility rise with each higher level of allowable risk along with higher realized total returns. Importantly, all mean-CVaR optimized portfolios provide improved risk/return profiles, each outperforming the benchmark portfolio showing positive alphas along with higher corresponding Sharpe ratios of around 0.65.

EXHIBIT 10

Next, we add step 3 into our process by incorporating the signal derived from our Markov-Switching risk model that identifies the current market state as being in either a high-risk or low-risk environment. We estimate our model under the expanding window approach with daily data beginning in February 1, 1996 with our first regime risk estimate occurring in January 3, 2000. If the risk model output suggests that the current environment is low-risk (high-risk) measured as less (more) than a 50% likelihood of being in a high-risk state, then we implement a risk-on (risk-off) strategy and optimize portfolio weights allowing for a CVaR risk target of 7% (3%), respectively. As before, expected return assumptions are the same for each state, and the optimization techniques are the same as those used in the constant target CVaR process, above. The only difference being that we now allow the target portfolio to change its risk profile to either risk-on (7% CVaR) or risk-off (3% CVaR) to reflect our dynamic forecast of market risk.

In this approach, we simply use the same rebalancing conditions as the constant CVaR process (i.e. same rebalancing dates and the constant 21-day rebalancing period). On rebalancing days, we choose the target CVaR for the upcoming 21-day period based on the prior day's market risk regime signal. Results show that incorporating the two-state market risk forecast meaningfully improves results over the benchmark and the constant CVaR approach. With this risk-on/risk-off framework, we are better able to capture a meaningful part of the upside that markets have to offer while also reducing the downside. This approach represents considerable improvement over the rebalanced, static benchmark and also the various static levels of CVaR. As evidence, consider that for the risk on/off model the Sharpe ratio rises to 0.68 while the maximum drawdown is now 19.73%, about half that of the benchmark. We note that this max drawdown is equivalent to that calculated under the 3% CVaR portfolio as calculated earlier but now captures much of the upside afforded by the risk-on days.⁶

Exhibit 10B summarizes our results by plotting the risk-return relationships for the various portfolios. Overall, results reflect the view that our CVaR tail risk framework offers a highly relevant risk measurement approach for investors. All CVaR related portfolios dominate the rebalanced static benchmark. Even the 3%

⁶ Our conclusions are unaffected when back-testing other rebalancing definitions.

CVaR, our lowest static risk portfolio, offered an appreciable excess return with far less risk than the benchmark. Adding a regime-based risk-on/off dynamic process enhances the performance even further. The regime-based optimization process outperforms the constant risk target allocations with significantly improved return/maximum drawdown ratios. Overall, our high frequency signal-triggered rebalancing framework offers a high degree of sensitivity of portfolio performance to market risk regime changes. That is, our flexible approach offers meaningful improvement in portfolio performance. Note that we discuss the bounded portfolio process, below. We offer these approaches as examples to allow readers a robust comparison of how various regime based strategies might perform over time. Finally, with each of our approaches, we are sensitive to keeping transactions costs associated with a high frequency of signals low, an issue facing many such dynamic frameworks.

Exhibit 11 shows the corresponding portfolio exposures of our five asset classes over time from our unbounded risk-on/risk-off regime-based approach that rebalances every 21 days. The exhibit shows the wide variation in asset allocation weightings generated by this approach. Many will consider the unbounded model presented here too demanding as it dictates dramatic shifts in portfolio asset allocation over time.

EXHIBIT 11

We appreciate the concern associated with such wide swings in asset weightings associated with portfolio rebalancing. To mitigate the wide swings in asset allocation associated with our unbounded approach, we next impose constraints on the range of allowable portfolio weights. Specifically, we test a constrained portfolio which allows the range of portfolio weights to vary only as much as that indicated in Exhibit 5. This bounded portfolio framework we test follows that of the 21-day rebalancing approach discussed above. Results show that performance versus a rebalanced static-weighted benchmark can benefit from our dynamic risk modeling framework even when imposing target portfolio constraints as typically done by many investors. As shown in Exhibits 10A and 10B, the bounded model offers considerable

improvement in both risk and return versus the rebalanced static-benchmark. Exhibit 12 shows the corresponding portfolio exposures over time for each of the five asset classes associated with the bounded risk-on/risk-off regime-based approach. As expected, it differs markedly from Exhibit 11. We note that during the global financial crisis of 2008-2009, given the minimum allowable allocation to risky assets, the model is unable to consistently achieve the desired 7% CVaR associated with a risk-off regime. This simply means that we are not always able to obtain the portfolio risk limits imposed when using a constrained approach with sizable minimum allowable allocations to risky assets.

EXHIBIT 12

Exhibit 13 shows the total cumulative returns to the various rebalancing approaches: benchmark, static allocations with 3%, 5% and 7% constant target CVaR, and a regime-based allocation that switches between 3% and 7% target CVaR under the same rebalancing conditions as the static allocations. This Exhibit offers visual evidence that our regime-based risk framework offers investors a meaningful approach to portfolio construction in the presence of fluctuating market risk.

Exhibit 13

Conclusions

We propose a dynamic portfolio construction model that accounts for the reality of heavy tails and dynamic return correlations as witnessed in markets. The powerful framework behind our portfolio construction is a dynamic process that integrates high-frequency information to capture the time-varying risks of asset classes within the investor's portfolio. We use our dynamic risk information to adjust optimal asset allocation across time and market states using only information known at the time of model implementation. We find that ongoing monitoring of markets using our market

risk barometer and corresponding asset allocation framework offers investors the promising opportunity to improve portfolio performance in challenging market environments.

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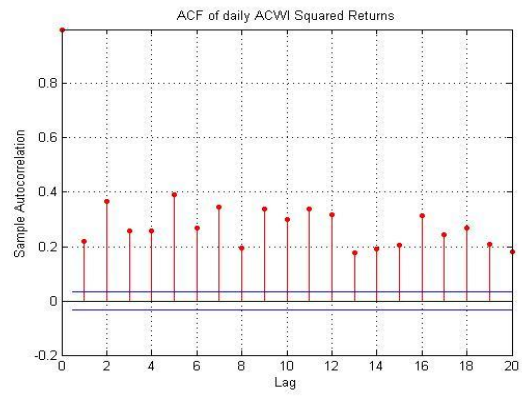
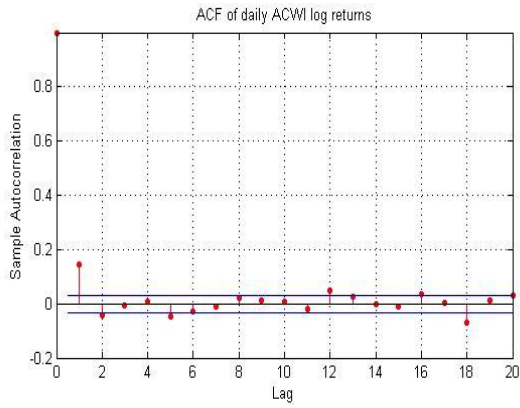
EXHIBIT 1 Asset Classes, Indices and Summary Statistics of Daily Returns

Period from February 1, 1996 to October 10, 2011

Index	Global Equity MSCI ACWI	Commodities SPGSCI	Real Estate DW REITs	High Yield MLHY	Investment Grade Barclay Agg
Mean	0.02%	0.03%	0.06%	0.03%	0.02%
Std Deviation	1.04%	1.50%	1.90%	0.29%	0.26%
Median	0.07%	0.03%	0.03%	0.05%	0.03%
Min	-7.10%	-8.76%	-19.76%	-4.73%	-2.04%
Max	9.31%	7.48%	18.98%	2.78%	1.71%
1st Percentile	-2.98%	-4.08%	-6.30%	-0.87%	-0.64%
99th Percentile	2.68%	3.62%	6.24%	0.73%	0.66%
5th Percentile	-1.61%	-2.40%	-2.32%	-0.36%	-0.40%
95th Percentile	1.53%	2.37%	2.16%	0.36%	0.41%
10th Percentile	-1.09%	-1.70%	-1.33%	-0.21%	-0.27%
90th Percentile	1.07%	1.82%	1.39%	0.24%	0.32%
Skewness	-0.24	-0.16	0.44	-2.74	-0.27
Kurtosis	7.50	2.38	21.21	42.55	3.11

EXHIBIT 2 Autocorrelations of Daily Returns and Squared Returns – ACWI

Panel A. Log Daily Returns



Panel B. Standardized Log Residual Returns

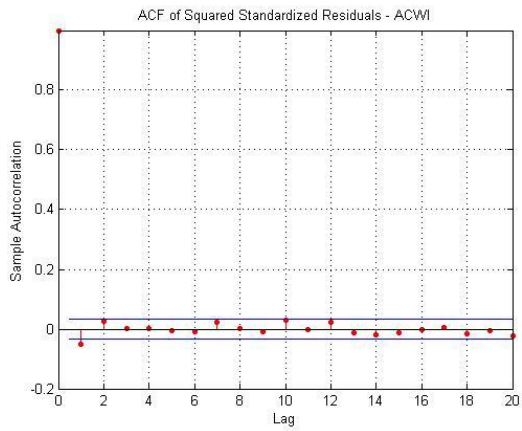
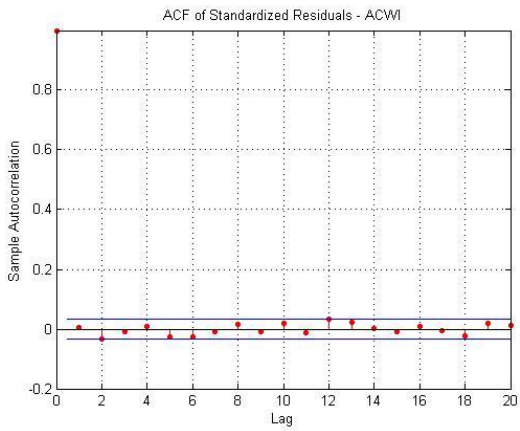


EXHIBIT 3 ACWI Lower and Upper Tail Fit

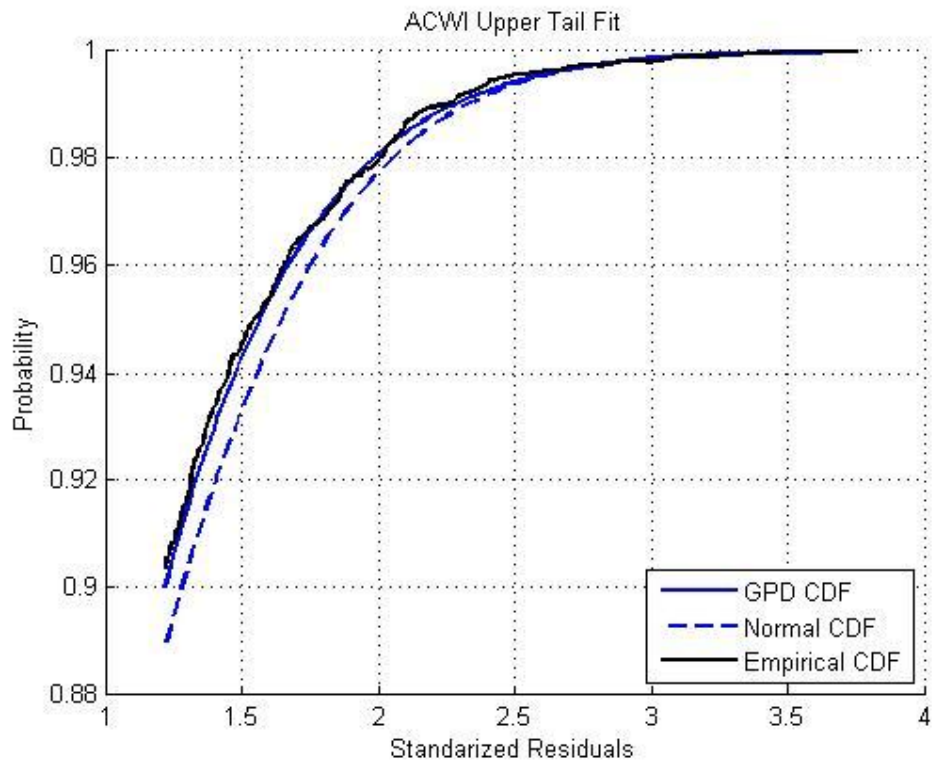
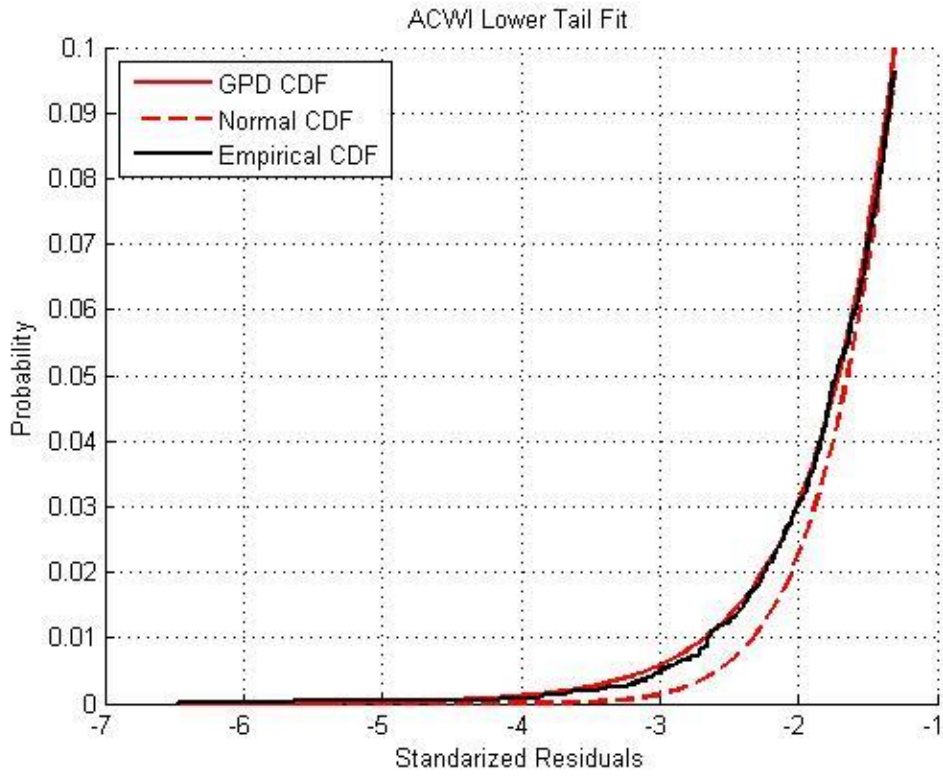


EXHIBIT 4 Scatter Plot of ACWI vs GSCI Log Daily Returns

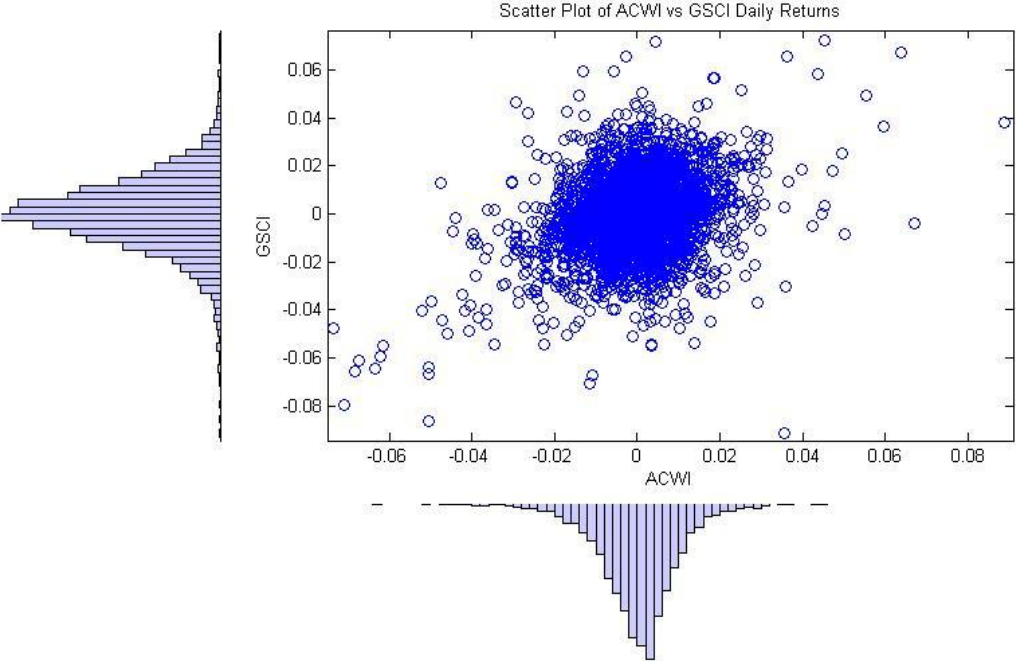


EXHIBIT 5 Portfolio Assumptions

Index	Global Equity (MSCI ACWI)	Commodities (SPGSCI)	Real Estate (DW REITs)	High Yield (MLHY)	Investment Grade (Barclay Agg)
Policy Allocation	45%	10%	10%	15%	20%
Portfolio Bounds	30 - 70%	5 - 15%	5 - 15%	7 - 23%	10 - 40%
Expected Returns	7%	6.5%	7%	6%	4%

EXHIBIT 6 Historical 21-day Forward-Looking 95% CVaR(Log)

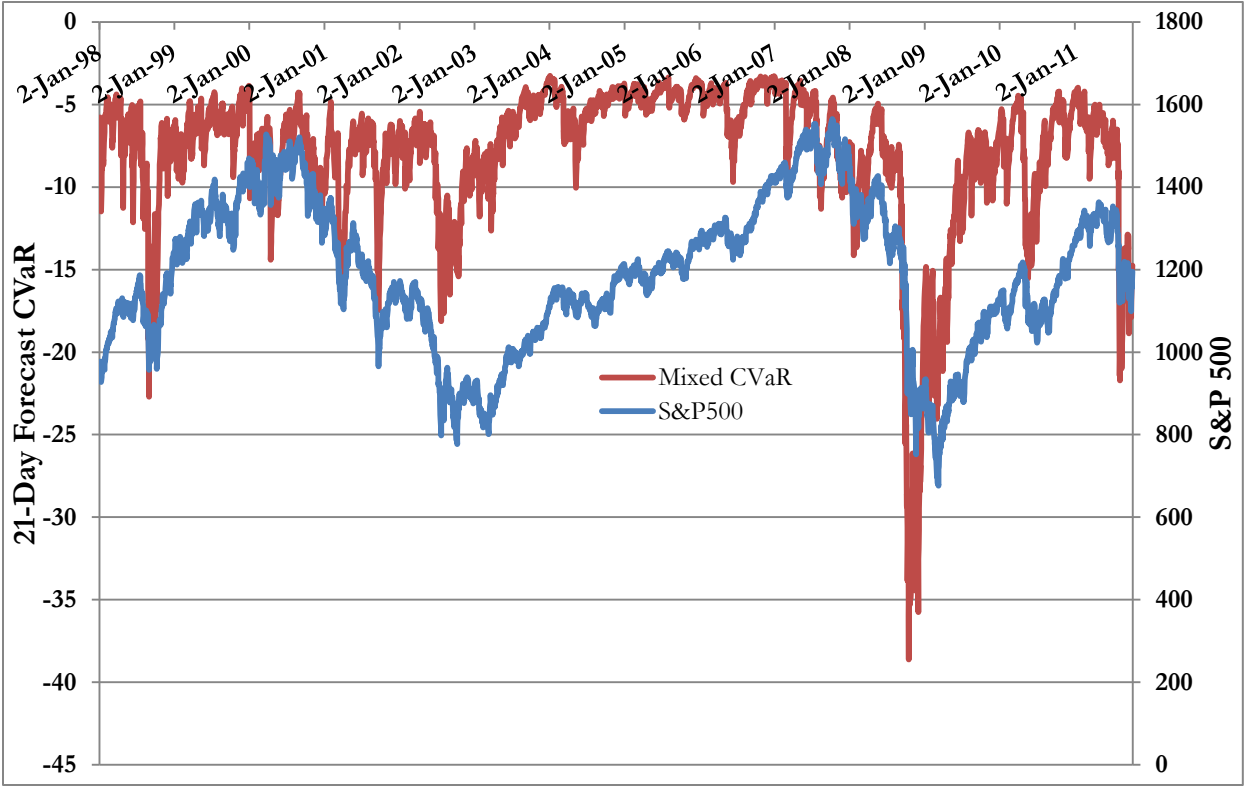


EXHIBIT 7 Markov-Switching Model Perfect Insights Estimation

	Regime 1			Regime 2 ("event")		
	Persistence	Mu	Sigma	Persistence	Mu	Sigma
Market Downside Risk	99.14%	-6.12%	1.63%	96.29%	-14.22%	5.69%

EXHIBIT 8 Expanding Window Approach 1/3/2000 – 10/10/2011

(A) Probability of the Event Regime

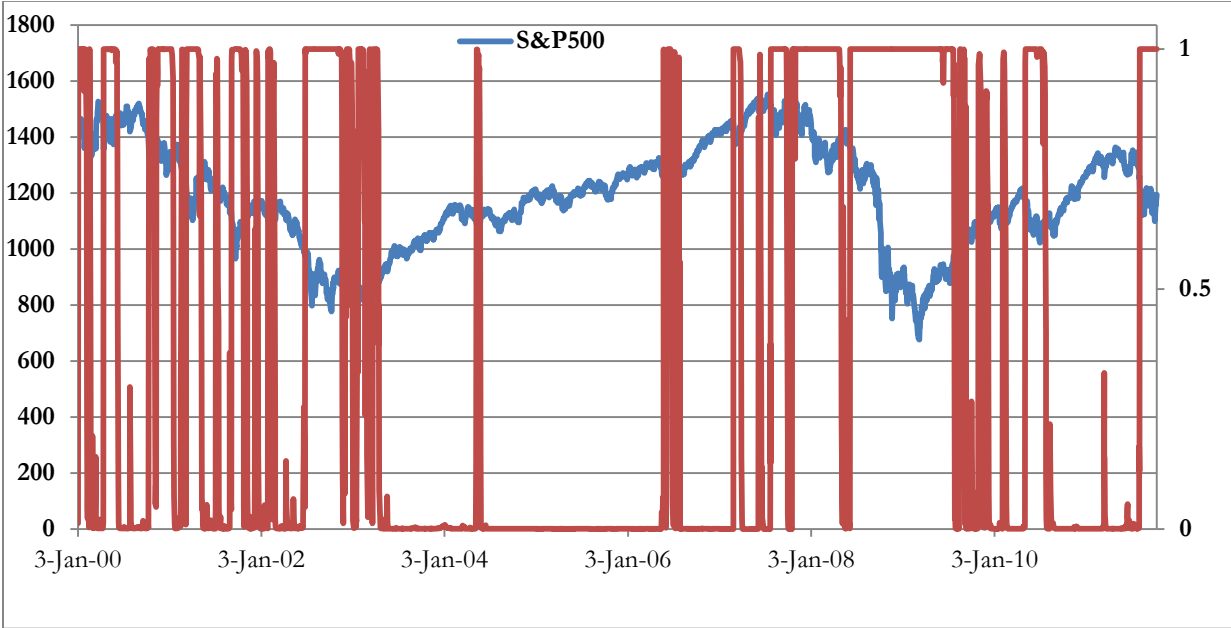


EXHIBIT 8 Expanding Window Approach 1/3/2000 – 10/10/2011

(B) Event Regime Periods

Start	4-Jan-00	9-Feb-00	18-Feb-00	14-Apr-00	11-Oct-00	10-Nov-00	20-Feb-01	9-Mar-01	6-Jul-01	6-Sep-01
End	8-Feb-00	10-Feb-00	23-Feb-00	9-Jun-00	6-Nov-00	18-Jan-01	6-Mar-01	9-May-01	17-Jul-01	24-Oct-01
Trading Days in Between	24	1	2	38	18	45	10	42	7	31
Start	29-Oct-01	10-Dec-01	13-Dec-01	29-Jan-02	19-Feb-02	26-Jun-02	26-Nov-02	5-Dec-02	9-Dec-02	27-Dec-02
End	8-Nov-01	11-Dec-01	21-Dec-01	13-Feb-02	26-Feb-02	21-Nov-02	27-Nov-02	6-Dec-02	26-Dec-02	2-Jan-03
Trading Days in Between	8	1	6	11	5	104	1	1	12	3
Start	21-Jan-03	24-Jan-03	24-Feb-03	10-Mar-03	24-Mar-03	10-May-04	19-May-06	2-Jun-06	6-Jun-06	6-Jul-06
End	23-Jan-03	21-Feb-03	28-Feb-03	19-Mar-03	15-Apr-03	25-May-04	23-May-06	5-Jun-06	30-Jun-06	7-Jul-06
Trading Days in Between	2	19	4	7	16	11	2	1	18	1
Start	21-Jul-06	25-Jul-06	27-Feb-07	7-Jun-07	12-Jun-07	26-Jul-07	19-Oct-07	9-May-08	6-Jun-08	27-Jul-09
End	24-Jul-06	26-Jul-06	30-Mar-07	8-Jun-07	15-Jun-07	1-Oct-07	2-May-08	12-May-08	24-Jul-09	28-Jul-09
Trading Days in Between	1	1	23	1	3	46	134	1	285	1
Start	17-Aug-09	3-Sep-09	9-Sep-09	30-Oct-09	27-Nov-09	4-Feb-10	4-May-10	18-Mar-11	4-Aug-11	
End	2-Sep-09	4-Sep-09	10-Sep-09	12-Nov-09	9-Dec-09	17-Feb-10	26-Jul-10	21-Mar-11	10-Oct-11	
Trading Days in Between	12	1	1	9	8	8	57	1	46	

EXHIBIT 9 Summary Statistics of Daily Returns 1/3/2000 – 10/10/2011

	Global Equity	Commodities	Real Estate	High Yield	Investment Grade
Index	MSCI ACWI	SPGSCI	DW REITs	MLHY	Barclay Agg
Mean	0.01%	-0.02%	0.05%	0.00%	0.03%
Std Deviation	1.59%	1.89%	3.32%	0.47%	0.30%
Median	0.08%	0.03%	0.03%	0.03%	0.04%
Min	-7.10%	-8.76%	-19.76%	-4.73%	-2.04%
Max	9.31%	7.48%	18.98%	2.78%	1.71%
1st Percentile	-4.90%	-5.35%	-9.27%	-1.62%	-0.70%
99th Percentile	4.41%	4.93%	11.12%	1.12%	0.79%
5th Percentile	-2.50%	-3.13%	-5.31%	-0.74%	-0.47%
95th Percentile	2.31%	2.81%	5.00%	0.64%	0.47%
10th Percentile	-1.75%	-2.26%	-3.00%	-0.46%	-0.34%
90th Percentile	1.66%	2.12%	3.17%	0.44%	0.37%
Skewness	-0.14	-0.28	0.33	-1.45	-0.18
Kurtosis	6.50	5.13	9.66	17.79	6.53

(A) Event Days

	MSCI ACWI	SPGSCI	DW REITs	MLHY	Barclay Agg
Mean	0.01%	0.03%	0.07%	0.03%	0.03%
Std Deviation	1.13%	1.62%	2.20%	0.33%	0.26%
Median	0.07%	0.05%	0.09%	0.05%	0.03%
Min	-7.10%	-8.76%	-19.76%	-4.73%	-2.04%
Max	9.31%	7.48%	18.98%	2.78%	1.71%
1st Percentile	-3.55%	-4.37%	-7.37%	-0.96%	-0.64%
99th Percentile	2.90%	3.84%	7.33%	0.84%	0.66%
5th Percentile	-1.74%	-2.62%	-2.88%	-0.42%	-0.41%
95th Percentile	1.61%	2.54%	2.61%	0.42%	0.42%
10th Percentile	-1.21%	-1.86%	-1.71%	-0.25%	-0.28%
90th Percentile	1.14%	2.00%	1.65%	0.28%	0.33%
Skewness	-0.21	-0.19	0.38	-2.51	-0.21
Kurtosis	9.86	5.02	18.63	38.30	6.13

(B) Full Period

EXHIBIT 10

Panel A Unbounded Optimization

	Info Ratio	Alpha	Tracking Error	Sharpe Ratio	Return / MaxDD	MaxDD	Worst 21day loss	Annual Return	Annual Volatility	Hit Ratio
Benchmark				0.24	0.11	42.83%	15.99%	4.84%	14.26%	
No leverage, no shorting										
Constant Target CVaR										
3%	0.23	2.91%	12.56%	0.67	0.39	19.73%	13.00%	7.75%	8.32%	49.65%
4%	0.48	4.25%	8.78%	0.67	0.33	27.62%	15.52%	9.09%	10.56%	53.90%
5%	0.65	4.81%	7.36%	0.63	0.29	33.22%	16.34%	9.65%	12.45%	56.03%
6%	0.78	5.47%	7.05%	0.61	0.27	37.71%	16.97%	10.31%	14.11%	65.96%
7%	0.79	5.68%	7.20%	0.59	0.26	40.90%	17.53%	10.52%	15.29%	63.83%
Regime Based Target CVaR (Same Rebalancing Conditions)										
3%, 7%	0.44	4.90%	11.09%	0.68	0.49	19.73%	14.84%	9.74%	11.39%	60.28%

Panel B Bounded Optimization

	Info Ratio	Alpha	Tracking Error	Sharpe Ratio	Return / MaxDD	MaxDD	Worst 21day loss	Annual Return	Annual Volatility	Hit Ratio
Bound										
Regime Based Target CVaR (Same Rebalancing Conditions)										
3%, 7%	0.37	1.40%	3.74%	0.37	0.18	35.36%	13.20%	6.24%	12.54%	59.57%

Exhibit 10B Portfolio Risk & Return Comparison

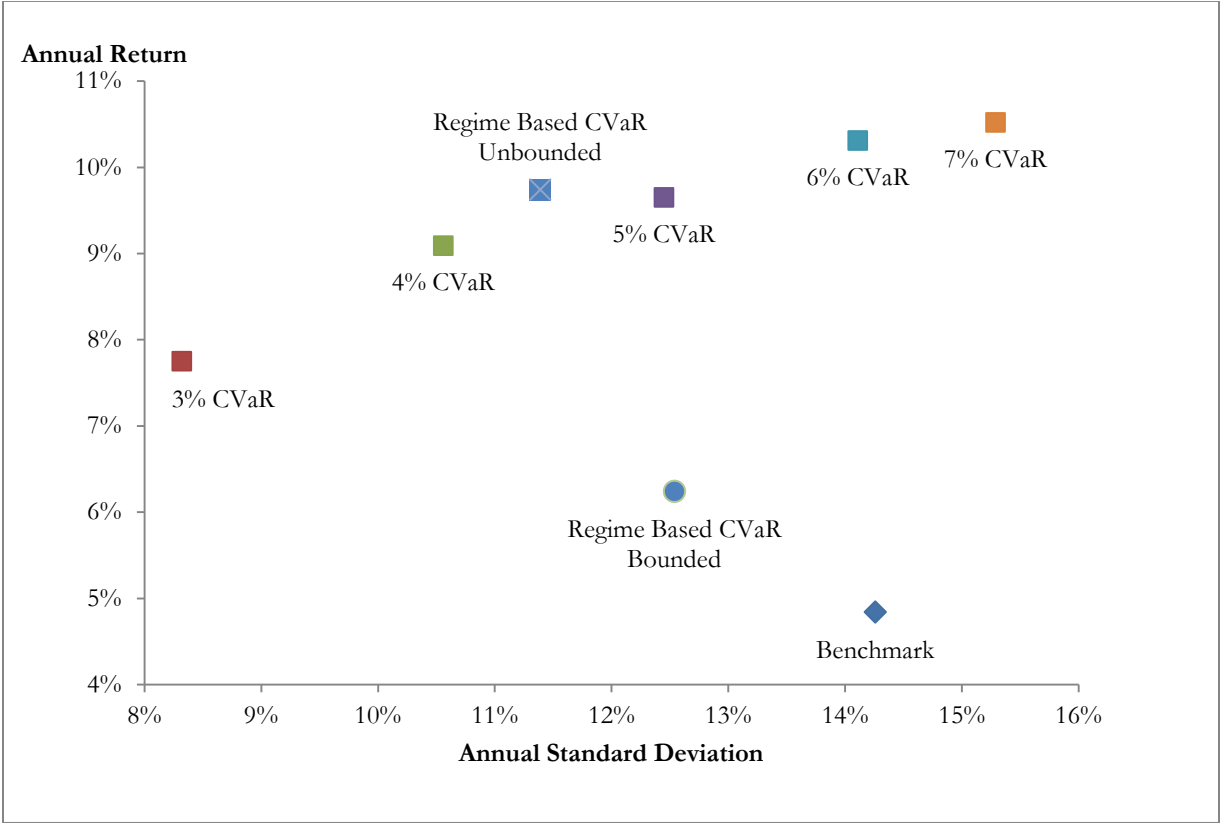


EXHIBIT 11 Optimal Portfolio Weights: Regime Based (3%, 7% Target CVaR)

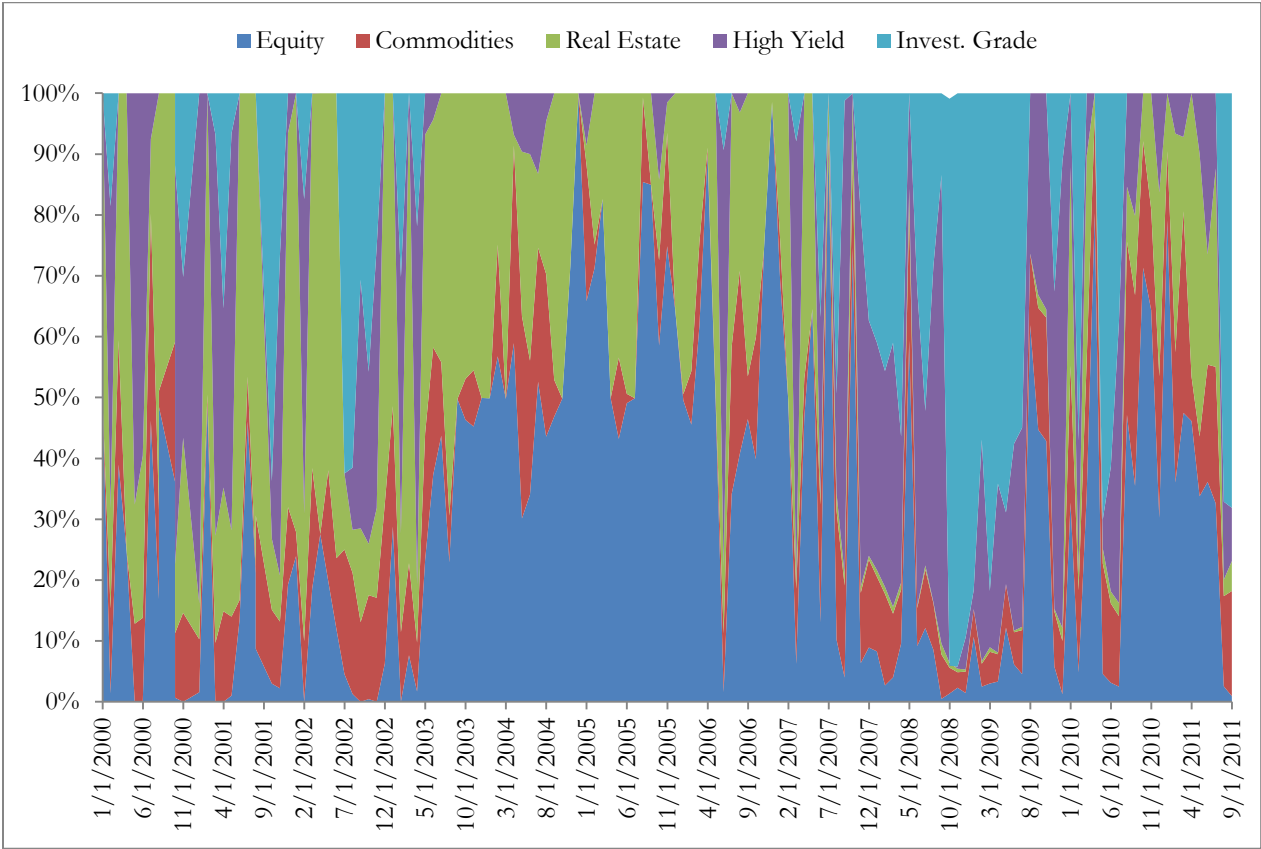


EXHIBIT 12 Optimal Portfolio Weights: Bounded Portfolio

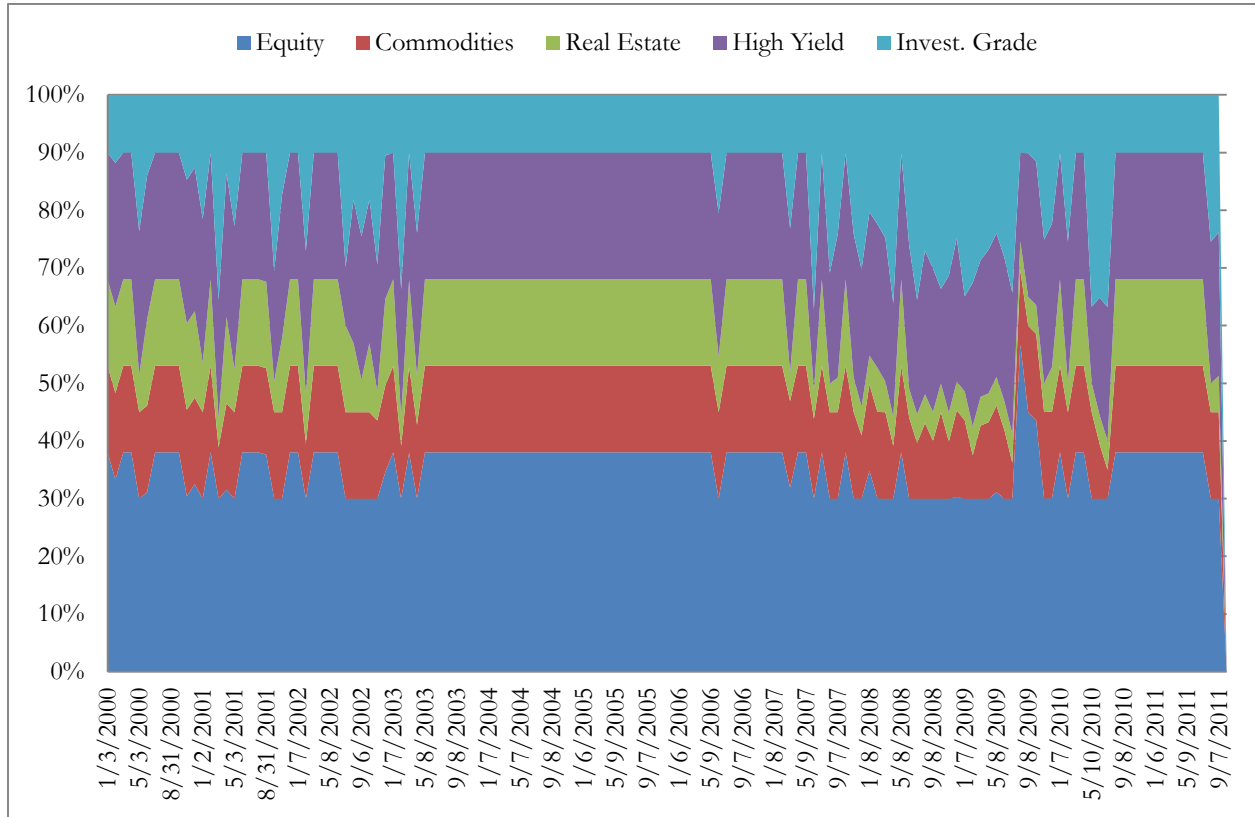


EXHIBIT 13 Cumulative Portfolio Returns

